



## Estimating Normalizing Constant (Partition Function, Free Energy)

**Task:** given an unnormalized probability density  $\pi \propto e^{-V}$ , estimate its normalizing constant (a.k.a. partition function)  $Z = \int_{\mathbb{R}^d} e^{-V(x)} dx$  or free energy  $F = -\log Z$ .

As a crucial problem in Bayesian statistics, statistical mechanics, and machine learning, it is challenging in high dimensions or when  $\pi$  is multimodal.

**Importance sampling:** with a prior  $\mu = \frac{1}{Z_\mu} e^{-U}$ , we have the equality  $\frac{Z_\pi}{Z_\mu} = \frac{1}{Z_\mu} \int e^{-V} dx = \mathbb{E}_\mu \frac{e^{-V}}{e^{-U}}$ . Hence the ratio can be estimated by sampling from  $\mu$ . However, this estimator suffers from high variance due to the mismatch between  $\mu$  and  $\pi$ .

## Annealing for Addressing Multimodality

**Annealing:** construct a sequence of intermediate distributions that bridge the target and the prior distributions. This idea motivates several popular methods:

- In statistics: path sampling, **annealed importance sampling**, sequential Monte Carlo, etc.
- In thermodynamics: thermodynamic integration, **Jarzynski equality**, etc.

**Contributions:** we aim to establish a rigorous non-asymptotic analysis of estimators based on JE and AIS, while introducing minimal assumptions on the target distribution. We also propose a new algorithm based on reverse diffusion samplers (RDS) to tackle a potential shortcoming of AIS.

## Wasserstein Distance, Metric Derivative, and Action

For probability measures  $\mu, \nu$  on  $\mathbb{R}^d$ , the **Wasserstein-2 distance** is defined as  $W_2(\mu, \nu) = \inf_{\gamma \in \Pi(\mu, \nu)} \left( \int \|x - y\|^2 \gamma(dx, dy) \right)^{\frac{1}{2}}$ , where  $\Pi(\mu, \nu)$  is the set of all couplings of  $(\mu, \nu)$ .

A vector field  $v = (v_t : \mathbb{R}^d \rightarrow \mathbb{R}^d)_{t \in [a, b]}$  on  $\mathbb{R}^d$  generates a curve of probability measures  $\rho = (\rho_t)_{t \in [a, b]}$  if the **continuity equation**  $\partial_t \rho_t + \nabla \cdot (\rho_t v_t) = 0$ ,  $t \in [a, b]$  holds.

The **metric derivative** of  $\rho$  at  $t \in [a, b]$  is defined as  $|\dot{\rho}|_t := \lim_{\delta \rightarrow 0} \frac{W_2(\rho_{t+\delta}, \rho_t)}{|\delta|}$ , which can be interpreted as the “speed” of this curve. If  $|\dot{\rho}|_t$  exists and is finite for a.e.  $t \in [a, b]$ , we say that  $\rho$  is **absolutely continuous (AC)**. Its **action** is defined as  $\int_a^b |\dot{\rho}|_t^2 dt$ , which is a key property characterizing the effectiveness of a curve in annealed sampling.

### ■ Lemma (Relationship between Metric Derivative and Continuity Equation [1])

For an AC curve of probability measures  $(\rho_t)_{t \in [a, b]}$ , any vector field  $(v_t)_{t \in [a, b]}$  that generates  $(\rho_t)_{t \in [a, b]}$  satisfies  $|\dot{\rho}|_t \leq \|v_t\|_{L^2(\rho_t)}$  for a.e.  $t \in [a, b]$ . Moreover, there exists a unique vector field  $(v_t^*)_{t \in [a, b]}$  generating  $(\rho_t)_{t \in [a, b]}$  that satisfies  $|\dot{\rho}|_t = \|v_t^*\|_{L^2(\rho_t)}$  for a.e.  $t \in [a, b]$ .

## Problem Setting

**Criterion:** given an accuracy threshold  $\varepsilon$ , study the oracle complexity required to obtain an estimator  $\hat{Z}$  of  $Z$  such that  $\Pr \left( \left| \frac{\hat{Z}}{Z} - 1 \right| \leq \varepsilon \right) \geq \frac{3}{4}$ . Note that the probability can be boosted to any  $1 - \zeta$  using the median trick.

**Annealing curve:** we define a curve of probability measures  $\left( \pi_\theta = \frac{1}{Z_\theta} e^{-V_\theta} \right)_{\theta \in [0, 1]}$  from a prior distribution to the target distribution.  $Z_1 = Z$  is what we need to estimate.

- **Assump. 1:** the potential  $[0, 1] \times \mathbb{R}^d \ni (\theta, x) \mapsto V_\theta(x) \in \mathbb{R}$  is jointly  $C^1$ , and the curve  $(\pi_\theta)_{\theta \in [0, 1]}$  is AC with finite action  $\mathcal{A} := \int_0^1 |\dot{\pi}|_\theta^2 d\theta$ .
- **Assump. 2:**  $V$  is  $\beta$ -smooth, and there exists  $x_*$ , with  $\|x_*\| =: R \lesssim \frac{1}{\sqrt{\beta}}$  s.t.  $\nabla V(x_*) = 0$ . Let  $m := \sqrt{\mathbb{E}_\pi \|\cdot\|^2} < +\infty$ .

## Analysis of the Jarzynski Equality (JE)

We introduce a reparameterized curve  $(\tilde{\pi}_t = \pi_{\frac{t}{T}})_{t \in [0, T]}$  for some large  $T$  to be determined later. **Annealed Langevin diffusion (ALD):**

$$dX_t = \nabla \log \tilde{\pi}_t(X_t) dt + \sqrt{2} dB_t, \quad t \in [0, T]; \quad X_0 \sim \tilde{\pi}_0.$$

### ► Jarzynski Equality (JE) [5]

Let  $\mathbb{P}^\rightarrow$  be the path measure of ALD. Then the following relation between the work functional  $W$  and free energy difference  $\Delta F$  holds:

$$\mathbb{E}_{\mathbb{P}^\rightarrow} e^{-W} = e^{-\Delta F}, \quad \text{where } W(X) := \frac{1}{T} \int_0^T \partial_\theta V_\theta|_{\theta=\frac{t}{T}}(X_t) dt, \quad \text{and } \Delta F := -\log \frac{Z_1}{Z_0}.$$

### ■ Theorem (Convergence Guarantee of JE)

$\hat{Z} := Z_0 e^{-W(X)}$  with  $X \sim \mathbb{P}^\rightarrow$  is an unbiased estimator of  $Z = Z_0 e^{-\Delta F}$ . Under Assump. 1, it suffices to choose  $T = \frac{32\mathcal{A}}{\varepsilon^2}$  to obtain  $\Pr \left( \left| \frac{\hat{Z}}{Z} - 1 \right| \leq \varepsilon \right) \geq \frac{3}{4}$ .

## Analysis of the Annealed Importance Sampling (AIS)

### ► Annealed Importance Sampling (AIS) Equality [6]

Suppose we have probability distributions  $\pi_\ell = \frac{1}{Z_\ell} f_\ell$ ,  $\ell \in \llbracket 0, M \rrbracket$  and transition kernels  $F_\ell(x, \cdot)$ ,  $\ell \in \llbracket 1, M \rrbracket$ , and assume that each  $\pi_\ell$  is an invariant distribution of  $F_\ell$ ,  $\ell \in \llbracket 1, M \rrbracket$ . Define the path measure  $\mathbb{P}^\rightarrow(x_{0:M}) = \pi_0(x_0) \prod_{\ell=1}^M F_\ell(x_{\ell-1}, x_\ell)$ . Then we have

$$\mathbb{E}_{\mathbb{P}^\rightarrow} e^{-W} = e^{-\Delta F}, \quad \text{where } W(x_{0:M}) := \log \prod_{\ell=0}^{M-1} \frac{f_\ell(x_\ell)}{f_{\ell+1}(x_\ell)} \quad \text{and } \Delta F := -\log \frac{Z_M}{Z_0}.$$

For non-asymptotic analysis, we focus on the **geometric interpolation**:

$$\pi_\theta = \frac{1}{Z_\theta} f_\theta = \frac{1}{Z_\theta} \exp \left( -V - \frac{\lambda(\theta)}{2} \|\cdot\|^2 \right), \quad \theta \in [0, 1]: \lambda(0) = 2\beta \searrow \lambda(1) = 0.$$

Introduce discrete time points  $0 = \theta_0 < \theta_1 < \dots < \theta_M = 1$ , and define  $F_\ell$  as running LD targeting  $\pi_{\theta_\ell}$  for time  $T_\ell$ . In practice, we approximate this by running one step of **annealed Langevin Monte Carlo (ALMC)** using the exponential integrator discretization scheme with step size  $T_\ell$ .

### ■ Theorem (Convergence Guarantee of AIS)

Under Assumps. 1 and 2, consider the annealing schedule  $\lambda(\theta) = 2\beta(1 - \theta)^r$  for some  $1 \leq r \lesssim 1$ . We use  $\mathcal{A}_r$  to denote the action of  $(\pi_\theta)_{\theta \in [0, 1]}$ . Then the oracle complexity for obtaining an estimate  $\hat{Z}$  that satisfies the criterion  $\Pr \left( \left| \frac{\hat{Z}}{Z} - 1 \right| \leq \varepsilon \right) \geq \frac{3}{4}$  is  $\tilde{\mathcal{O}} \left( \frac{d^{\frac{3}{2}}}{\varepsilon^2} \vee \frac{m\beta\mathcal{A}_r^{\frac{1}{2}}}{\varepsilon^2} \vee \frac{d\beta^2\mathcal{A}_r^2}{\varepsilon^4} \right)$ .

## Reverse Diffusion Sampler (RDS)

The choice of the curve  $(\pi_\theta)_{\theta \in [0, 1]}$  is crucial for the complexity of JE & AIS. The geometric interpolation is widely used due to the availability of the scores of  $\pi_\theta$ . However, for general target distributions, the action of the curve can be large:

### ■ Lemma (Exponential Lower Bound on the Action of Geometric Annealing)

Consider  $\pi = \frac{1}{2} \mathcal{N}(0, 1) + \frac{1}{2} \mathcal{N}(m, 1)$  on  $\mathbb{R}$  for some large  $m \gtrsim 1$ , whose potential is  $\frac{m^2}{2}$ -smooth. Under the setting in AIS, define  $\pi_\theta(x) \propto \pi(x) e^{-\frac{\lambda(\theta)}{2} x^2}$ ,  $\theta \in [0, 1]$ , where  $\lambda(\theta) = m^2(1 - \theta)^r$  for some  $1 \leq r \lesssim 1$ . Then, the action of the curve  $(\pi_\theta)_{\theta \in [0, 1]}$  is  $\mathcal{A}_r \gtrsim m^4 e^{\frac{m^2}{40}}$ .

**Reverse diffusion samplers (RDS):** a series of multimodal samplers inspired by diffusion models. The OU process  $dY_t = -Y_t dt + \sqrt{2} dB_t$ ,  $t \in [0, T]$ ;  $Y_0 \sim \pi$  transforms any target distribution  $\pi$  into  $\phi := \mathcal{N}(0, I)$  as  $T \rightarrow \infty$ . Let  $Y_t \sim \bar{\pi}_t$ . The time-reversal  $(Y_t^\leftarrow := Y_{T-t} \sim \bar{\pi}_{T-t})_{t \in [0, T]}$  satisfies the SDE  $dY_t^\leftarrow = (Y_t^\leftarrow + 2\nabla \log \bar{\pi}_{T-t}(Y_t^\leftarrow)) dt + \sqrt{2} dW_t$ ,  $t \in [0, T]$ ;  $Y_0^\leftarrow \sim \bar{\pi}_T (\approx \phi)$ . We propose leveraging the curve along the OU process for normalizing constant estimation. The following proposition supports this idea:

### ■ Proposition (Polynomial Upper Bound of the Action of the OU curve)

Let  $\bar{\pi}_t$  be the law of  $Y_t$  in the OU process initialized from  $Y_0 \sim \pi \propto e^{-V}$ , where  $V$  is  $\beta$ -smooth and let  $m^2 := \mathbb{E}_\pi \|\cdot\|^2 < \infty$ . Then,  $\int_0^\infty |\dot{\bar{\pi}}|_t^2 dt \leq d\beta + m^2$ .

### ■ Theorem (a Framework for Normalizing Constant Estimation via RDS)

Assume a total time duration  $T$ , an early stopping time  $\delta \geq 0$ , and discrete time points  $0 = t_0 < t_1 < \dots < t_N = T - \delta \leq T$ . For  $t \in [0, T - \delta)$ , let  $t_-$  denote  $t_k$  if  $t \in [t_k, t_{k+1})$ . Let  $s_- \approx \nabla \log \bar{\pi}_t$  be a score estimator, and  $\phi = \mathcal{N}(0, I)$ . Consider the following two SDEs on  $[0, T - \delta]$  representing the sampling trajectory and the time-reversed OU process, respectively:

$$\begin{aligned} \mathbb{Q}^\dagger: \quad dX_t &= (X_t + 2s_{T-t}(X_{t_-})) dt + \sqrt{2} dB_t, & X_0 &\sim \phi; \\ \mathbb{Q}: \quad dX_t &= (X_t + 2\nabla \log \bar{\pi}_{T-t}(X_t)) dt + \sqrt{2} dB_t, & X_0 &\sim \bar{\pi}_T. \end{aligned}$$

Let  $\hat{Z} := e^{-W(X)}$ ,  $X \sim \mathbb{Q}^\dagger$  be the estimator of  $Z$ , where  $X \mapsto W(X)$  is defined as

$$\log \phi(X_0) + V(X_{T-\delta}) + (T - \delta) d + \int_0^{T-\delta} \left( \|s_{T-t}(X_{t_-})\|^2 dt + \sqrt{2} \langle s_{T-t}(X_{t_-}), dB_t \rangle \right).$$

Then, to ensure  $\hat{Z}$  satisfies  $\Pr \left( \left| \frac{\hat{Z}}{Z} - 1 \right| \leq \varepsilon \right) \geq \frac{3}{4}$ , it suffices that  $\text{KL}(\mathbb{Q} \parallel \mathbb{Q}^\dagger) \lesssim \varepsilon^2$ ,  $\text{TV}(\pi, \bar{\pi}_\delta) \lesssim \varepsilon$ . We can use results in [3, 4, 2, 7] to derive the total complexity.

## References

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